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ASYMPTOTIC EXPANSION FOR THE DISTRIBUTION OF NONLINEAR RENEWAL PROCESS

Abstract

In present paper the asymptotic expansion of high order for distribution of process of nonlinear reconstruction.

I. Introduction. Let $\xi_n, n \geq 1$ be independent identically distributed random variables determined on some probability space (Ω, \mathcal{F}, P) and let $f_a(t), t > 0, a > 0$ be a family of nonlinear (non-random) Borellian functions.

Put:

$$S_0 = 0, \quad S_n = \sum_{k=1}^n \xi_k, \quad n \geq 1 \quad \text{is the}$$

$$\tau_a = \inf \{n \geq 1 : S_n > f_a(n)\}$$

is the first exit moment beyond the nonlinear boundary $f_a(t)$.

Asymptotic expansion for the distribution of markovian moment τ_a has been studied in paper [1,2]. The first two terms for the asymptotic expansions on the boundary $f_a(t)$, and on the distributions of the random variables ξ_1 have been obtained in [2]. In this paper, the highest order asymptotic expansion for the distribution of τ_a is studied.

2. Conditions and notations. Assume, that for each $a > 0$ the function $f_a(t) > 0$ and it is continuously differentiable for $t > 0$, where

$$\sup_{a,t} f'_a(t) < M = \text{const}. \tag{1}$$

We can easily verify that for the family of boundaries $f_a(t) = K(t+a)^\beta + c$, where $K > 0, c \geq 0, 0 \leq \beta < 1$, the condition (1) is fulfilled with $M = K \cdot \beta$. Further, assume that:

$$P(\xi_1 \geq M) = 1$$

and

$$\rho_m = E|\xi - M\xi|^m < \infty, \quad m = \overline{3, r}. \tag{2}$$

Denote:

$$C_n = \frac{f_a(n) - n\mu}{\sigma\sqrt{n}},$$

where $\mu = E\xi_1$ and $\sigma^2 = D\xi_1$.

To derive the highest order asymptotic expansions we assume that the characteristic function $\psi_\xi(\lambda)$ of r.v. ξ_1 must satisfy the following condition:

$$\overline{\lim}_{|\lambda| \rightarrow \infty} |\psi_\xi(\lambda)| < 1. \tag{3}$$

Remark that the condition (3) is fulfilled for any nonsingular distribution of r.v. ξ_1 [3], and also for any strongly non-latticed distribution [4].

We also denote:

$$\Phi(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{u^2}{2}} du$$

$$\varphi(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}.$$

3. The basic result. Let the above enumerated conditions (1), (2) and (3) be fulfilled. Besides, let $C_n = O(1)$ for $n = n(a) \rightarrow \infty$.

Then:

$$P(\tau_a \leq n) = \Phi(-C_n) - \varphi(C_n) \cdot \sum_{k=3}^r n^{\frac{k}{2}+1} R_k(C_n) + o\left(n^{\frac{r}{2}+1}\right),$$

where $R_k = R_n(x)$, $x \in (R)$ is a polynomial with real coefficients depending only on $\rho_1, \rho_2, \dots, \rho_r$, but not on n and z (or on other characteristics of the distribution ξ_1). See [3, p.607] for the construction of the polynomial $R_k(x)$.

4. Proof of the basic result.

We need the following

Lemma. Let the conditions (1) and (2) be fulfilled. Then

$$P(\tau_a \leq n) = P(S_n > fa(n))$$

holds.

Proof. It is clear that

$$\{\omega : \tau_a > n\} = \{\omega : \max_{k \leq n} (S_k - fa(k)) \leq 0\}. \quad (4)$$

We have:

$$\max_{1 \leq k \leq n} (S_k - fa(k)) = \max_{1 \leq k \leq n} \left[\sum_{i=1}^k (\xi_i - fa(i) - fa(i-1)) + fa(0) \right]. \quad (5)$$

It follows from the condition (2), that for each i with probability 1

$$\xi_i - (fa(i) - fa(i-1)) \geq 0.$$

Therefore, it follows from (5), that

$$\max_{1 \leq k \leq n} (S_k - fa(k)) = \sum_{i=1}^n (\xi_i - fa(i) - fa(i-1)) + fa(0) = S_n - fa(n).$$

Now the statement of the lemma follows from (4) and (5).

Now prove the statement of the theorem.

Denote

$$F_n(x) = P\left(\frac{S_n - n\mu}{\sigma\sqrt{n}} \leq x\right).$$

By virtue of the lemma we have

$$P(\tau \leq n) = 1 - F_n(C_n). \quad (6)$$

Under the assumptions concerning the finiteness of the moment $p_r < \infty$ for some r , it holds the Edgeworths expression [3].

$$F_n(x) = \Phi(x) + \varphi(x) \sum_{i=3}^r n^{\frac{i}{2}+1} R_i(x) + o\left(n^{\frac{r}{2}+1}\right) \quad (7)$$

uniformly in x .

Considering that $\Phi(x) + \Phi(-x) = 1$, we obtain from (6) and (7) the affirmation of the theorem.

Note that the statement of the theorem for $r = 3$ is contained in paper [2], where random variables with values of different signs have been considered.

References

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